Variational inference for Dirichlet process mixtures

Mikolaj Kasprzak (<u>mjkasprz@mit.edu</u>) Ali Ramadhan (<u>alir@mit.edu</u>) 6.435 Bayesian Modeling & Inference Lecture 14 April 7, 2022

Blei & Jordan (2006), Bayesian Analysis 1(1), pp. 121-144.

Plan for today

- 1. What is a Dirichlet Process and DP mixture and why do we care?
- 2. What is Variational Inference and why is it useful?
- 3. Why do we need Variational Inference for DP mixtures?
- 4. How can we go about constructing a VI algorithm for DP mixtures?
- 5. Demo: Dirichlet process mixture model for cluster assignment
- 6. Comparison between VI and Gibbs sampling for DP mixtures.
- 7. Empirical evaluation and example applications.

What is a Dirichlet Process?

- <u>Parameters:</u> $\alpha > 0, G_0$ probability distribution
- Nature of the object:
 - $G \sim DP(\alpha, G_0)$ is a random discrete probability distribution



<u>Question</u>: Is DP a useful object? What are its potential applications?

What is a DP mixture?

- 1. <u>Parameters:</u> $\alpha > 0, G_0$ probability distribution
- 2. <u>The construction:</u>
 - 1. Draw $V_1, V_2, \ldots | \alpha \sim \text{Beta}(1, \alpha)$ i.i.d
 - 2. Draw $\eta_1^*, \eta_2^*, \dots | G_0 \sim G_0$ i.i.d.
 - 3. For the nth data point:
 - Draw $Z_n | \{v_1, v_2, \dots\} \sim \text{Categorical}(\pi(v))$
 - Draw $X_n | z_n \sim p(x_n | \eta_{z_n}^*)$
- 3. Additional assumptions the paper makes:
 - $X_n | \{z_n, \eta_1^*, \eta_2^*, ...\}$ comes from an exponential family
 - G_0 is the corresponding conjugate prior.



Exponential families

- 1. <u>Density</u>: $p(x|\boldsymbol{\eta}) = h(x) \exp\left[\boldsymbol{\eta}^T \mathbf{T}(x) \mathbf{a}(\boldsymbol{\eta})\right]$, for some functions $h, \mathbf{T}, \mathbf{a}$.
- 2. <u>Conjugate prior</u>: $p_{\text{prior}}(\boldsymbol{\eta}|\boldsymbol{\chi},\nu) \propto \exp\left[\boldsymbol{\eta}^T \boldsymbol{\chi} \nu \mathbf{a}(\boldsymbol{\eta})\right]$
- 3. <u>Examples of exponential families:</u> normal, exponential, gamma, beta, Dirichlet, chi-squared, Bernoulli, categorical, Poisson, geometric, binomial (with fixed number of trials), multinomial (with fixed number of trials).
- 4. In our case:

$$p(x_n|z_n,\eta_1^*,\eta_2^*...) = \prod_{i=1}^{\infty} \left[h(x_n) \exp\left[(\eta_i^*)^T x_n - a(\eta_i^*) \right] \right]^{1[z_n=i]}$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

$$p_{\text{prior}}(\eta^*|\lambda) = h(\eta^*) \exp\left[\lambda_1^T \eta^* - \lambda_2 a(\eta^*) - a(\lambda) \right]$$

• Draw $X_n | z_n \sim p(x_n | \eta_{z_n}^*)$

Variational inference

- 1. <u>Suppose</u> you want to approximate the posterior distribution of latent variables: $p_{posterior}(\boldsymbol{w}|\boldsymbol{x},\theta) = \exp\left[\log p_{X,W}(\boldsymbol{x},\boldsymbol{w}|\theta) - \log p_X(\boldsymbol{x}|\theta)\right]$
- 2. <u>Choose</u> a family of variational distributions: $\{q_{\nu}(w) : \nu \in \{\text{space of variational parameters}\}\}$ 3. Minimize the KL divergence (with respect to the variational parameter)
- $D(q_{\nu}(\boldsymbol{w}) || p(\boldsymbol{w} | \boldsymbol{x}, \theta)) = \mathbb{E}_{\boldsymbol{W} \sim q_{\nu}} \log q_{\nu}(\boldsymbol{W}) \mathbb{E}_{\boldsymbol{W} \sim q_{\nu}} \log p_{X,W}(\boldsymbol{x}, \boldsymbol{W} | \theta) + \log p_{X}(\boldsymbol{x} | \theta)$ 4. Equivalently, maximize the ELBO (with respect to the variational parameter)

$$\text{ELBO} = \mathbb{E}_{\boldsymbol{W} \sim q_{\nu}} \log p_{X,W}(\boldsymbol{x}, \boldsymbol{W} | \theta) - \mathbb{E}_{\boldsymbol{W} \sim q_{\nu}} \log q_{\nu}(\boldsymbol{W})$$

5. <u>Mean-field VI</u>: For an M-dimensional latent vector **W**, with conditionals following an exponential-family distribution $p(w_i|\boldsymbol{w}_{-i}, \boldsymbol{x}, \theta)$, choose

$$q_{\nu}(\boldsymbol{w}) = \prod_{i=1}^{M} \exp\left(v_i^T w_i - a(w_i)\right)$$

6. <u>Question:</u> Is Variational Inference useful? Do we need VI for DP mixtures? Can we do mean-field VI as described above, in this case?

Mean-field VI for DP mixtures

- 1. Latent variables: stick lengths, the atoms and cluster assignments: $W = \{V, \eta^*, Z\}$
- 2. <u>Hyperparameters</u>: scaling parameter and the parameter of G_0 : $\theta = \{\alpha, \lambda\}$
- 3. <u>How</u> to choose a variational family to approximate an infinite-dimensional random object depending on infinite sets $V = \{V_1, V_2, ...\}$ and $\eta^* = \{\eta_1^*, \eta_2^*, ...\}$?
- 4. Truncate! Stop breaking the stick! Fix some T and let the variational family satisfy $q(v_T = 1) = 1$. This means that the mixture proportions $\pi_t(v) = 0$, for all t > T.
- 5. The variational family is:

$$q_{\nu}(\boldsymbol{v}, \boldsymbol{v}^{*}, \boldsymbol{z}) = \prod_{t=1}^{T-1} q_{\gamma_{t}}(v_{t}) \prod_{t=1}^{T} q_{\tau_{t}}(\eta_{t}^{*}) \prod_{n=1}^{N} q_{\phi_{n}}(z_{n})$$

where $q_{\gamma_t}(v_t)$ are beta, $q_{\tau_t}(\eta_t^*)$ are exponential family and $q_{\phi_n}(z_n)$ are categorical.

6. The free variational parameter: $\boldsymbol{\nu} = \{\gamma_1, \dots, \gamma_{T-1}, \tau_1, \dots, \tau_T, \phi_1, \dots, \phi_N\}$

A few questions

- 1. Why did we choose the stick-breaking representation of the Dirichlet Process to do VI?
- 2. Can we use the truncation trick and still call ourselves non-parametric Bayesians?
- 3. Does your answer to the previous question change if you a) truncate your generative model or b) truncate the variational approximation?
- 4. In general, what do we lose by truncating?
- 5. Is truncation the only reasonable approach? Can you think of other ways of bringing

the DP mixture down to a finite-dimensional world?

Minimizing KL aka maximizing ELBO ELBO = $\mathbb{E}_{W \sim q_{\nu}} \log p_{X,W}(W, x|\theta) - \mathbb{E}_{W \sim q_{\nu}} \log q_{\nu}(W)$ = $\mathbb{E}_q \left[\log p(V|\alpha) \right] + \mathbb{E}_q \left[\log p(\eta^*|\lambda) \right] + \sum_{n=1}^N \left(\mathbb{E}_q [\log p(Z_n|V)] + \mathbb{E}_q [\log p(x_n|Z_n)] \right)$

 $-\mathbb{E}_q[\log q(\boldsymbol{V}, \boldsymbol{\eta}^*, \boldsymbol{Z})]$

Use truncation to express the red term as a finite sum and apply coordinate ascent!

Calculating the predictive posterior:

$$p(X_{N+1}|\boldsymbol{X}, \alpha, \lambda) = \int \left(\sum_{t=1}^{\infty} \pi_t(\boldsymbol{v}) p(x_{N+1}|\boldsymbol{\eta}_t^*)\right) dP(\boldsymbol{v}, \boldsymbol{\eta}^*|\boldsymbol{x}, \lambda, \alpha)$$
$$\approx \sum_{t=1}^{T} \mathbb{E}_q[\pi_t(\boldsymbol{V})] \mathbb{E}_q\left[p(x_{N+1}|\boldsymbol{\eta}_t^*)\right]$$

Question: Is the approximate predictive posterior easier to calculate than the true one? Why?

Demo: Let's do some variational inference!





https://raw.githubusercontent.com/ali-ramadhan/random-jupyter-note s/master/Bavesian/clusters.mp4

Using Dirichlet process mixture models

- Pyro: <u>https://pyro.ai/examples/dirichlet_process_mixture.html</u>
- PyMC3:<u>https://docs.pymc.io/en/v3/pymc-examples/examples/mixture_models/</u> <u>dp_mix.html</u>
- sklearn.mixture.DPGMM:<u>https://scikit-learn.org/0.15/modules/generated/sklear</u> <u>n.mixture.DPGMM.html</u>
- Turing.jl: <u>https://turing.ml/dev/tutorials/06-infinite-mixture-model/</u>
- Comparison: <u>https://luiarthur.github.io/TuringBnpBenchmarks/dpsbgmm</u>

Empirical comparison: Variational inference vs. Gibbs sampling

- VI can be faster.
- VI can use the Evidence lower bound (ELBO) can to assess convergence.
- VI optimization can fall into local maxima.

40

20

0

-20

0

- VI only produces an approximation.
- No theory for evaluating VI disadvantages so the authors turn to empirical comparisons.

Problem: Dirichlet process mixture model for 100 data points sampled from a 2D DP mixture of Gaussians with diagonal covariance.



Empirical comparison: Variational inference vs. Gibbs sampling

- *Collapsed Gibbs*: marginalize over one or more variables when sampling for some other variable.
 - Integrate out G and η . Just sample cluster assignment c for each data point.
- Blocked Gibbs: group two or more variables together and samples from their joint distribution conditioned on all other variables.
 - Use a truncated Dirichlet process (TDP).



Empirical comparison: Variational inference vs. Gibbs sampling

• Testing on held-out data: treat each data point as the 101st data point and compute its conditional probability.

Dim	Mean held out log probability (Std err)		
	Variational	Collapsed Gibbs	Truncated Gibbs
5	-147.96(4.12)	-148.08(3.93)	-147.93(3.88)
10	-266.59(7.69)	-266.29(7.64)	-265.89(7.66)
20	-494.12(7.31)	-492.32(7.54)	-491.96(7.59)
30	-721.55 (8.18)	-720.05(7.92)	-720.02(7.96)
40	-943.39(10.65)	-941.04(10.15)	$-940.71 \ (10.23)$
50	-1151.01 (15.23)	-1148.51 (14.78)	-1147.48(14.55)





Image analysis

Analyze 5,000 images reduced to an 8x8 grid of average RGB values.



00

Conclusions

- 1. DP mixture models are useful for applications where the number of clusters (categories, topics, ...) may potentially grow without a bound.
- 2. In order to do inference with DP mixture models in practice we need to approximate the posterior with some tractable, finite-dimensional distribution.
- 3. The authors describe a computationally efficient mean-field variational inference algorithm for DP mixture models, which outputs such an approximation.
- 4. They use the stick-breaking construction of DP and truncate this procedure at some fixed point in order to be in a finite-dimensional world.
- 5. Variational inference for DP mixture models can be faster and scale better than MCMC methods (Gibbs sampling).